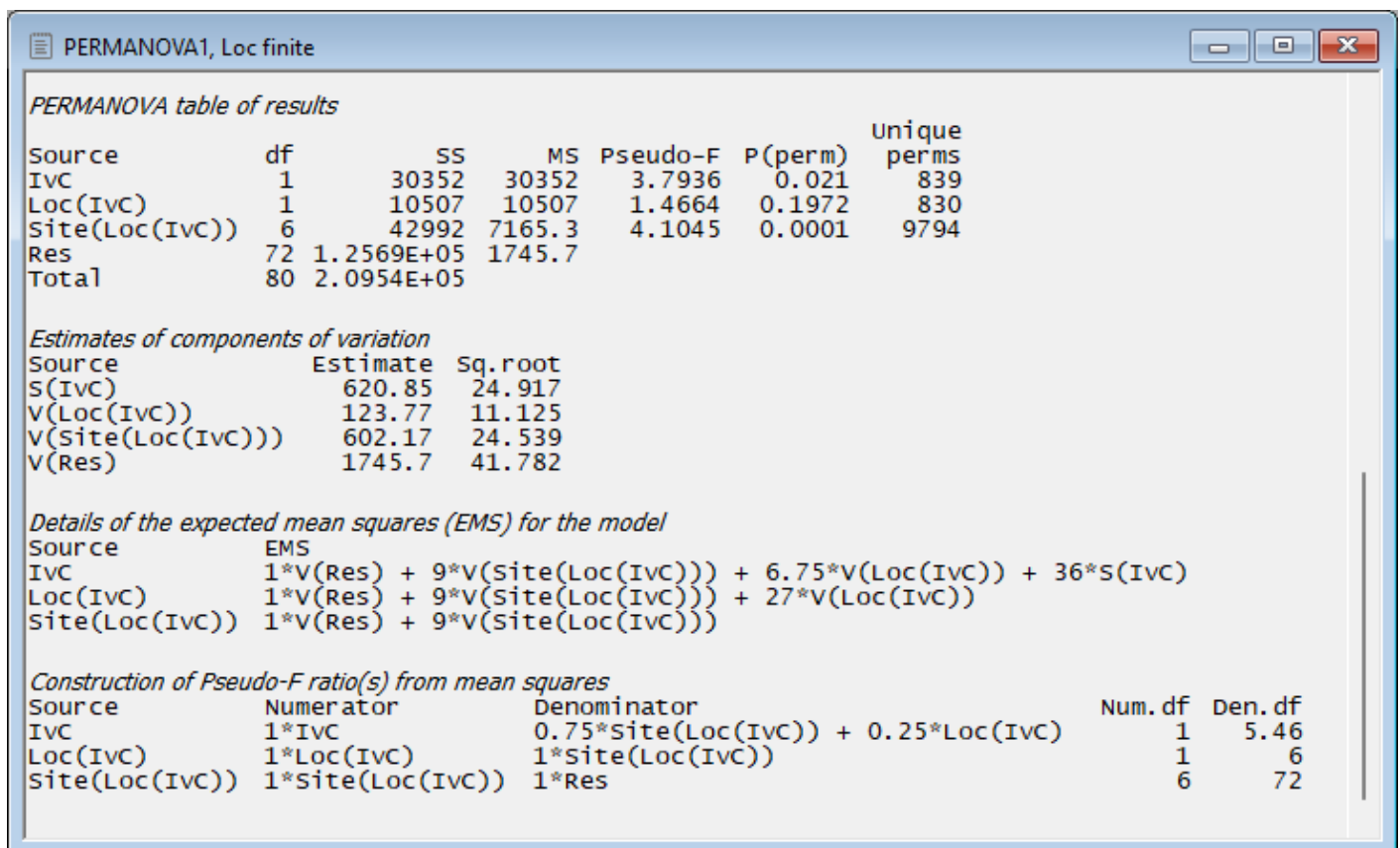


7.5 Broader implications for detecting impact

Comparison of results treating 'Locations' as random

Historical wisdom for such a design would have treated 'Locations' as a random factor ([Underwood \(1992\)](#) , [Glasby \(1997\)](#)). It is quite instructive to consider what the results of this analysis might have been had we done this, instead of treating the 'Locations' factor as finite. Below are the two output files:

- treating 'Loc' as a *finite factor* with a sampling fraction of $2/8 (= 1/4)$ for the controls ('PERMANOVA1')

A screenshot of a software window titled "PERMANOVA1, Loc finite". The window displays three sections of statistical output. The first section, "PERMANOVA table of results", is a table with 7 columns: Source, df, SS, MS, Pseudo-F, P(perms), and Unique perms. The second section, "Estimates of components of variation", is a table with 3 columns: Source, Estimate, and Sq.root. The third section, "Details of the expected mean squares (EMS) for the model", is a table with 2 columns: Source and EMS. Below this is another section, "Construction of Pseudo-F ratio(s) from mean squares", which is a table with 5 columns: Source, Numerator, Denominator, Num. df, and Den. df.

PERMANOVA table of results						
Source	df	SS	MS	Pseudo-F	P(perms)	Unique perms
IvC	1	30352	30352	3.7936	0.021	839
Loc(IvC)	1	10507	10507	1.4664	0.1972	830
Site(Loc(IvC))	6	42992	7165.3	4.1045	0.0001	9794
Res	72	1.2569E+05	1745.7			
Total	80	2.0954E+05				

Estimates of components of variation		
Source	Estimate	Sq.root
S(IvC)	620.85	24.917
V(Loc(IvC))	123.77	11.125
V(Site(Loc(IvC)))	602.17	24.539
V(Res)	1745.7	41.782

Details of the expected mean squares (EMS) for the model	
Source	EMS
IvC	$1 \cdot V(\text{Res}) + 9 \cdot V(\text{Site}(\text{Loc}(\text{IvC}))) + 6.75 \cdot V(\text{Loc}(\text{IvC})) + 36 \cdot S(\text{IvC})$
Loc(IvC)	$1 \cdot V(\text{Res}) + 9 \cdot V(\text{Site}(\text{Loc}(\text{IvC}))) + 27 \cdot V(\text{Loc}(\text{IvC}))$
Site(Loc(IvC))	$1 \cdot V(\text{Res}) + 9 \cdot V(\text{Site}(\text{Loc}(\text{IvC})))$

Construction of Pseudo-F ratio(s) from mean squares				
Source	Numerator	Denominator	Num. df	Den. df
IvC	$1 \cdot \text{IvC}$	$0.75 \cdot \text{Site}(\text{Loc}(\text{IvC})) + 0.25 \cdot \text{Loc}(\text{IvC})$	1	5.46
Loc(IvC)	$1 \cdot \text{Loc}(\text{IvC})$	$1 \cdot \text{Site}(\text{Loc}(\text{IvC}))$	1	6
Site(Loc(IvC))	$1 \cdot \text{Site}(\text{Loc}(\text{IvC}))$	$1 \cdot \text{Res}$	6	72

- treating 'Loc' as a *random factor* ('PERMANOVA2').

PERMANOVA2, Loc random

PERMANOVA table of results

Source	df	SS	MS	Pseudo-F	P(perm)	Unique perms
IvC	1	30352	30352	2.8886	0.3309	3
Loc(IvC)	1	10507	10507	1.4664	0.1997	830
Site(Loc(IvC))	6	42992	7165.3	4.1045	0.0001	9810
Res	72	1.2569E+05	1745.7			
Total	80	2.0954E+05				

Estimates of components of variation

Source	Estimate	Sq.root
S(IvC)	551.23	23.478
V(Loc(IvC))	123.77	11.125
V(Site(Loc(IvC)))	602.17	24.539
V(Res)	1745.7	41.782

Details of the expected mean squares (EMS) for the model

Source	EMS
IvC	$1 \cdot V(\text{Res}) + 9 \cdot V(\text{Site}(\text{Loc}(\text{IvC}))) + 27 \cdot V(\text{Loc}(\text{IvC})) + 36 \cdot S(\text{IvC})$
Loc(IvC)	$1 \cdot V(\text{Res}) + 9 \cdot V(\text{Site}(\text{Loc}(\text{IvC}))) + 27 \cdot V(\text{Loc}(\text{IvC}))$
Site(Loc(IvC))	$1 \cdot V(\text{Res}) + 9 \cdot V(\text{Site}(\text{Loc}(\text{IvC})))$

Construction of Pseudo-F ratio(s) from mean squares

Source	Numerator	Denominator	Num. df	Den. df
IvC	$1 \cdot \text{IvC}$	$1 \cdot \text{Loc}(\text{IvC})$	1	1
Loc(IvC)	$1 \cdot \text{Loc}(\text{IvC})$	$1 \cdot \text{Site}(\text{Loc}(\text{IvC}))$	1	6
Site(Loc(IvC))	$1 \cdot \text{Site}(\text{Loc}(\text{IvC}))$	$1 \cdot \text{Res}$	6	72

There are a number of things that are different between these two outputs (Table 7.2). Look specifically at the EMS for the 'IvC' term and, hence, the construction of the pseudo F ratio from mean squares and associated degrees of freedom for that test. These, in turn, obviously affect the observed value of the pseudo F statistic for the test of 'IvC' and its associated P value as well.

Table 7.2. Key essential differences in the PERMANOVA output when we treat 'Locations' as a finite population with 8 levels versus treating it as a random factor.

Point of difference	Treat 'Loc' as finite (8 levels)	Treat 'Loc' as random
Number of levels in the population for the factor of 'Location' (among Controls)	$B_1 = 8$	$B_1 = \infty$
Coefficient, K , on the 'Loc' variance component (Loc(IvC)) in the EMS for 'IvC'	$K = 6.75$	$K = 27$
Denominator in the construction of the pseudo F test statistic to test 'IvC'	$0.75 \cdot \text{MS}_{\text{Sites}} + 0.25 \cdot \text{MS}_{\text{Loc}}$	MS_{Loc}
Denominator degrees of freedom for the test of 'IvC'	$\text{df}_{\text{denom}} = 5.46$	$\text{df}_{\text{denom}} = 1$
pseudo F test-statistic for 'IvC'	$F = 3.7936$	$F = 2.8886$
P value for the test of 'IvC'	$P = 0.021$	$P = 0.331$ [¶]

An interesting thing to note is that the denominator that needs to be used for the construction of the pseudo F test-statistic for the test of 'IvC' when we treat 'Loc' as finite has to be constructed

as a *linear combination of mean squares*. This is also the essential reason that the denominator degrees of freedom for the test of 'IvC' in the finite-factor case is a non-integer value (i.e., $df_{\text{denom}} = 5.46$).

The implications of all of this are far-reaching, because the test of the 'IvC' term is definitely the most important test for the researcher in this particular study design. The specification of the 'Loc' factor as 'finite' has clearly provided *more power* for this key test of environmental impact in the present case. In general, we can expect that the power will increase whenever there is an increase in the denominator degrees of freedom (all else being equal).

Changes in the size of the inference space

To demonstrate the effect of changing the *size of the inference space* on the analysis, we can posit what the results for this study would look like - specifically for the test of the 'IvC' term in the model - if the number of 'Control' locations in the *population* (i.e., B_1), were larger (Table 7.3). We have already seen what would happen if we consider that this population is infinite (i.e., if we treat 'Locations' as a random factor). The table below looks at a progression of values for B_1 , corresponding to a gradation in the sampling fraction ($f = b_1/B_1$) from fixed ($f=1$) to random (where $B_1 = \infty$ so f is effectively zero), showing the concomitant change in the results.

Table 7.3. Effect of a change in the size of the inference space (sampling fraction) on construction of the F statistic and associated denominator degrees of freedom in a PERMANOVA test for the factor 'IvC'.

B_1	Sampling fraction (f)	Denominator for the test of 'IvC'	df_{denom}	F_{IvC}
∞ (random)	$\lim_{B_1 \rightarrow \infty} f = 0$	MS_{Loc}	1	2.889
100	1/50	$0.67 \cdot MS_{\text{Sites}} + 0.33 \cdot MS_{\text{Loc}}$	4.35	3.676
30	1/15	$0.69 \cdot MS_{\text{Sites}} + 0.31 \cdot MS_{\text{Loc}}$	4.57	3.699
20	1/10	$0.70 \cdot MS_{\text{Sites}} + 0.30 \cdot MS_{\text{Loc}}$	4.72	3.716
10	1/5	$0.73 \cdot MS_{\text{Sites}} + 0.27 \cdot MS_{\text{Loc}}$	5.21	3.767

B_1	Sampling fraction (f)	Denominator for the test of 'IvC'	$\frac{df}{denom}$	F_{IvC}
8	1/4	$0.75 \cdot MS_{\text{Sites}} + 0.25 \cdot MS_{\text{Loc}}$	5.46	3.794
4	1/2	$0.83 \cdot MS_{\text{Sites}} + 0.17 \cdot MS_{\text{Loc}}$	6.62	3.930
2	1	MS_{Sites}	6	4.235

Note that we are not changing anything about the number of levels *actually sampled*, here: i.e., for all of the lines in Table 7.3 above, we have the same $B_1 = 2$ sampled control locations.

A word of caution

We are not typically at liberty simply to choose whatever inference space we want. Ethical as well as scientific considerations may well come in to play here. It has to be recognised that the specific pseudo F ratio in every line of Table 7.3 actually corresponds to a test of a different null hypothesis. Each line presents a test with a different breadth of inference: the top line has the broadest scope, while the bottom line has the narrowest. Indeed, treating the control locations as fixed (bottom line in Table 7.3) might well give us more power, but it also severely limits our statistical inferences to just those particular locations in our study and to no others. It is clear that the inferences from the true study design, where we sampled 2 out of 8 control locations, apply to the whole of the Italian Mediterranean coastline that is spanned by those 8 potential locations, no less and no further. Hence, this excellent new tool permitting specification of finite factors, although it affords us greater flexibility and (potentially) power to test the terms of greatest interest, also comes with a special dose of responsibility. We need, as researchers, to articulate carefully the broader population, hence the scale and extent of the inferences that we shall draw from any study.

[¶]Note that the p -value here is being limited by the fact that there are only 3 possible permutations of the 3 locations across the 2 groups 'I' and 'C'. We do have the option to obtain a Monte Carlo approximation to the p -value when we run the PERMANOVA, by ticking the box in the PERMANOVA run dialog: (☒Do Monte Carlo tests). Assuming that each of the principal coordinate (PCO) axes representing the sample points in the chosen resemblance space (Bray-Curtis in this example) are asymptotically normal, then the PERMANOVA pseudo- F statistic is distributed as a ratio of two linear forms in chi-square under a true null hypothesis, from which we can take a random Monte Carlo draw (the linear forms are supplied by the eigenvalues of the PCO). The Monte Carlo approximate p -value for the test of the 'IvC' term for this example, treating 'Locations' as a random and not a finite factor, is $P \sim 0.03$.

Revision #73

Created 1 May 2025 21:16:43 by Marti

Updated 4 December 2025 04:33:58 by Marti